

TRANSIENT ANALYSIS OF TANDEM QUEUEING MODEL WITH NON-HOMOGENOUS POISSON BULK ARRIVALS HAVING STATE-DEPENDENT SERVICE RATES

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ABSTRACT

Queueing models provides the basic frame work for analysing the practical situations in production processes, communication networks, transportation systems, and machine repairing. In many of the communication systems the arrivals are time-dependent and can be characterised by a non-homogeneous Poisson process. In this paper we develop and analyze a two-node tandem queueing model with the assumption that the arrivals follows a non-homogeneous compound Poisson process, and the service rates of each node depends on the number of customers in the queue connected it. Using the difference - differential equations and a probability generating function of the number of customers in the queue and in the system is analysed. Explicit expression are derived for performance measures including average number of customers in each queue, the probability of emptiness of the system, the average waiting time in the queue and in the system, the throughput of the nodes, and the variance of the number of customers in the queue. Sensitivity analysis of the model with respect to the uniformly distributed batch size arrivals is also studied. It is observed that the system performance measures are highly influenced by non-homogeneous (time-dependent) arrivals rate and batch size distribution parameters. This model also includes several earlier queueing models as particular cases for specific values of the parameters. These models are useful in scheduling tele and wireless communication networks.

Keywords: Non- Homogeneous Poisson process, Bulk arrivals, State-dependent service rates, Tandem Queueing models.

1. INTRODUCTION :

Queueing models are essential for designing and monitoring of several communication systems. The queueing models provide the optimum operating policies of the several systems. They are essential for performance evaluation of a wide variety of systems in communication networks, production manufacturing process, data/voice transmission, nero physiological systems, cargo handling, and ATM scheduling. [1]. Starting from the first queueing model by A.K. Erlang [37] much work has been reported in literature regarding queueing models and their applications. These models are formed with suitable assumptions on arrival, service processes. In many of the practical situations the output from one queueing process serves as input to other i.e., the queues are connected in series. These types of queueing systems are called tandem

queueing systems. O' Brien [2], Jackson [3,5], Paul [4] have pioneered the tandem queueing models. Recently Yukuo Hayashida [6], Srinivasa Rao et al.[7], Yongjian et al.[8], Lieshout and Mandjes [9], Wai Kin Victor Chan [10], Che soong Kin et al. [11,12] and others have studied various tandem queueing models. In all these papers they considered that the arrival and service processes are independent. In many practical situations like communication networks the service rate is to be may dependent on the number of customers in the queue. This type of queueing models are called as queueing models with state-dependent service rates or queueing models with load dependent. To have accurate prediction of the performance measures of the system and to obtain optimal policies the load dependent queueing models are developed and analysed. [13-18]. In all these papers the authors assumed that the arrivals to the system come singly and can be characterised by a Poisson process.

In some practical situations like stored and forward communication networks the arrivals cannot be characterized by a Poisson process the arriving messages are converted into a random number of packets depending upon the size of the message. Similarly, in railway yards and ports the cargo handling is done in batches of random size. For analyzing these sort of situations the queueing models with bulk arrivals are developed. The bulk arrivals can be well characterised by a compound Poisson process. The compound Poisson process also includes Poisson process as a particular case. The bulk arrival queueing models were initiated by Erlang Solution of M/Ek/1 model Brockmeyer et al [19]. Later, several authors developed various queueing models for bulk arrivals. Recently, Madan et al [20] studied a single-server queue with batch arrivals having two types of heterogeneous service with a different service rates having general service time distribution. Chaudhry and Chang [21] developed a discrete time bulk-service queueing model known as the Geo/G^Y/1/N+B model. Juan [22] developed M/G^Y/1 a queueing model, with a discretised service time distribution. Schleyer and Furmans [23] presented an analytical method to calculate the waiting time distribution for the G/G/1 queueing system with batch arrivals. Ahmed [24] considered a multi-channel bi-level heterogeneous-server bulk arrival queueing system with an Erlangian service time. Chen A, Pollett P., LiJ., and Zhang H [25] studied a modified Markovian bulk arrival and bulk service queue incorporating state-dependent control. Charan Jeet Singh et al. (2011) studied a single-server bulk queueing system with state dependent rates and a second optional service. They used the supplementary variable technique to obtain the probability generating function. Dieter Claeys et al. [26] studied a threshold-based service system with batch arrivals and general service times. Arumuganathan et al. [27] studied two-node tandem communication network models with bulk arrivals using queueing theory. In all these models, the authors considered the arrivals to be homogeneous and follow compound Poisson process.

However, in queueing models connected with communication networks like Ethernet LAN, metropolitan area network (MAN) traffic, wider area network (WAN), and variable bit rate (VBR) traffic exhibit time-dependent arrival rates and cannot be modelled with homogeneous Poisson process or compound Poisson process [28-31]. The studies made by Crovella et al. [32], Murali Krishna et al. [33], Feldmann A. [34], Fischer et al. [35] also revealed that in a TCP connect communication network, the time between packet arrivals cannot be characterized by an exponential distribution. The observations made by Dinda et al. [36] have revealed that the traffic generated by many real world applications exhibits a high degree of burstiness (time varying arrival rate). Very little work has been reported regarding tandem queueing models with non homogeneous bulk arrivals. For this purpose, in this paper, we develop and analyse a two-node tandem queueing model with the assumption that the arrival process follows a non-homogeneous compound Poisson process having state-dependent service rates. The non-homogeneous compound Poisson process is capable of including a wide spectrum of Poisson process and can accommodate heterogeneous or burstiness and non smooth traffic arrivals.

Using difference - differential equations the joint probability generating function of the number of customers in each queue is derived. The system performance measures like the average number of customers in the queue and system, the average waiting time of the customers in the system and queue, the throughput of the node, etc., are derived. The sensitivity of the model with respect to the parameters is also studied by

assuming that the batch size of the arrivals follows a uniform distribution. A comparative study of the developed model with that of homogeneous and non homogenous (time dependent) arrivals is also discussed.

2. QUEUEING MODEL:

In this section, we consider two queues Q_1, Q_2 and service stations S_1, S_2 which are connected as series in order. It is assumed that the customers after getting service through first service station may join the second queue which is in series connected to S_1 . It is further assumed that the customers arrive to the first queue in batches of random size but dependent on time i.e., the actual number of customers in any arriving module is a random variable X with probability C_x . In other words, the arrival of customers follows non-homogeneous compound Poisson processes with a mean composite arrival rate $\lambda(t)$ having bulk size distributions C_x . The service completion in both the service stations follows Poisson processes with the parameters μ_1 and μ_2 for the first and second service stations. It is further assumed that the mean service rate in the service station is linearly dependent on the content of the queue connected to it. The queue discipline is first in, first out. A schematic diagram representing the queueing model is shown in figure 2.1

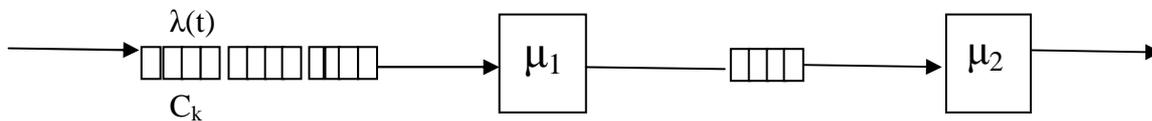


Figure 2.1 Schematic diagram of the queueing model

Let n_1 and n_2 denote the number of customers in first and second queues and let $P_{n_1, n_2}(t)$ be the probability that there are n_1 customers in the first queue and n_2 customers in the second queue at time t . The difference - differential equations governing the model are as follows:

$$\begin{aligned} \frac{\partial P_{n_1, n_2}(t)}{\partial t} &= -(\lambda(t) + n_1\mu_1 + n_2\mu_2)P_{n_1, n_2}(t) + \lambda(t) \sum_{k=1}^{n_1} P_{n_1-k, n_2}(t) C_k \\ &\quad + (n_1 + 1)\mu_1 P_{n_1+1, n_2-1}(t) + (n_2 + 1)\mu_2 P_{n_1, n_2+1}(t), \quad \text{for } n_1, n_2 > 0 \\ \frac{\partial P_{n_1, 0}(t)}{\partial t} &= -(\lambda(t) + n_1\mu_1)P_{n_1, 0}(t) + \lambda(t) \left[\sum_{k=1}^{n_1} P_{n_1-k, 0}(t) C_k \right] + \mu_2 P_{n_1+1}(t), \text{ for } n_1 > 0, n_2 = 0 \\ \frac{\partial P_{0, n_2}(t)}{\partial t} &= -(\lambda(t) + n_2\mu_2)P_{0, n_2}(t) + \mu_1 P_{1, n_2-1}(t) + (n_2 + 1)\mu_2 P_{0, n_2+1}(t), \text{ for } n_2 > 0, n_1 = 0 \\ \frac{\partial P_{0, 0}(t)}{\partial t} &= -\lambda(t)P_{0, 0}(t) + \mu_2 P_{0, 1}(t), \quad \text{for } n_1, n_2 = 0 \\ \frac{\partial P_{1, 0}(t)}{\partial t} &= -(\lambda(t) + \mu_1)P_{1, 0}(t) + \mu_2 P_{1, 1}(t) + \lambda(t)P_{0, 0}(t)C_1, \quad \text{for } n_1 = 1, n_2 = 0 \\ \frac{\partial P_{0, 1}(t)}{\partial t} &= -(\lambda(t) + \mu_2)P_{0, 1}(t) + \mu_1 P_{1, 0}(t) + 2\mu_2 P_{0, 2}(t), \quad \text{for } n_1 = 0, n_2 = 1 \end{aligned} \tag{1}$$

Let the joint probability generating function of $P_{n_1, n_2}(t)$ is be

$$P(Z_1, Z_2, t) = \sum_{n_1=0}^{\infty} \sum_{n_2=0}^{\infty} Z_1^{n_1} Z_2^{n_2} P_{n_1, n_2}(t) \tag{2}$$

Multiplying the equation (2.1) by Z_1, Z_2 and summing over all n_1 and n_2 , we obtain

$$\frac{\partial P(Z_1, Z_2, t)}{\partial t} = [-\mu_1(Z_1 - Z_2)] \frac{\partial P(Z_1, Z_2, t)}{\partial Z_1} + [\mu_2(Z_2 - 1)] \frac{\partial P(Z_1, Z_2, t)}{\partial Z_2} + P(Z_1, Z_2, t) [\lambda(t)(C(Z) - 1)]$$

where $C(Z) = \sum_{k=1}^{\infty} Z^k C_k$ is the probability generating function for the arriving batch size distribution of Q_1 .

Solving the equation (3) by the Lagrangian method, the auxiliary equations are

$$\frac{\partial t}{1} = \frac{\partial Z_1}{\mu_1(Z_1 - Z_2)} = \frac{\partial Z_2}{\mu_2(Z_2 - 1)} = \frac{\partial P(Z_1, Z_2, t)}{P(Z_1, Z_2, t) [\lambda(t)(C(Z) - 1)]}$$

Solving the first and third terms in equation (4), we obtain

$$a = (Z_2 - 1) e^{-\mu_2 t}$$

Solving the first and second terms in equation (4), we obtain

$$b = (Z_1 - 1) e^{-\mu_1 t} + a \mu_1 \frac{e^{(\mu_2 - \mu_1)t}}{(\mu_2 - \mu_1)}$$

Solving the first and fifth terms in equation (4), we obtain

$$c = P(Z_1, Z_2, t) \exp \left\{ \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i C_k \binom{k}{r} \binom{r}{i} b^{r-i} \left[\frac{\alpha \mu_1}{\mu_2 - \mu_1} \right]^i \left[\frac{-\lambda}{t\mu_2 + (r-i)\mu_1} + \frac{\alpha}{[t\mu_2 + (r-i)\mu_1]^2} \right] \right\}$$

where a, b and c are arbitrary constants derived using the initial conditions $P_{00}(0) = 1, P_{00}(t) = 0 \forall t > 0$

(5)

The general solution of (2.4) gives the probability generating function of the number of customers in the first queue and the number of customers in the second queue at time ‘t’ as

$$P(Z_1, Z_2, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i C_k \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^i \left[(Z_1 - 1) + \frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^{r-i} \left[\frac{1 - e^{-[t\mu_2 + (r-i)\mu_1]t}}{t\mu_2 + (r-i)\mu_1} + \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i C_k \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^i \left[(Z_1 - 1) + \frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^{r-i} \left[\frac{t[t\mu_2 + (r-i)\mu_1] - 1 + e^{-[t\mu_2 + (r-i)\mu_1]t}}{[t\mu_2 + (r-i)\mu_1]^2} \right] \right\}, |Z_1| < 1, |Z_2| < 1 \tag{6}$$

3. CHARACTERISTICS OF THE QUEUEING MODEL:

Expanding $P(Z_1, Z_2, t)$ given in equation (6) and collecting the constant terms, we obtain the probability that the system is empty to be

$$\begin{aligned}
 &P_{00}(t) \\
 &= \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} C_k \binom{k}{r} \binom{r}{i} \left[\frac{\mu_2^{r-i} \mu_1^i}{(\mu_2 - \mu_1)^r} \right] \left[\frac{1 - e^{-[i\mu_2 + (r-i)\mu_1]t}}{i\mu_2 + (r-i)\mu_1} \right] \right. \\
 &+ \alpha \left. \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} C_k \binom{k}{r} \binom{r}{i} \left[\frac{\mu_2^{r-i} \mu_1^i}{(\mu_2 - \mu_1)^r} \right] \left[\frac{t[i\mu_2 + (r-i)\mu_1] - 1 + e^{-[i\mu_2 + (r-i)\mu_1]t}}{[i\mu_2 + (r-i)\mu_1]^2} \right] \right\}
 \end{aligned} \tag{7}$$

Taking $Z_2 = 1$ in $P(Z_1, Z_2, t)$, we determine probability generating function for the first queue size to be

$$\begin{aligned}
 P(Z_1, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k C_k \binom{k}{r} (Z_1 - 1)^r \left[\frac{1 - e^{-r\mu_1 t}}{r\mu_1} \right] \right. \\
 \left. + \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k C_k \binom{k}{r} (Z_1 - 1)^r \left[\frac{tr\mu_1 - 1 + e^{-r\mu_1 t}}{[r\mu_1]^2} \right] \right\}
 \end{aligned} \tag{8}$$

By expanding $P(Z_1, t)$ and collecting the constant terms, we determine the probability that the first queue is empty to be

$$P_{00}(t) = \exp\{B_1(t)\} \text{ say}$$

$$\text{where } B_1(t) = \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k C_k \binom{k}{r} (-1)^r \left[\frac{1 - e^{-r\mu_1 t}}{r\mu_1} \right] + \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k C_k \binom{k}{r} (-1)^r \left[\frac{tr\mu_1 - 1 + e^{-r\mu_1 t}}{[r\mu_1]^2} \right]$$

(9)

The mean number of customers in the first queue is

$$L_1(t) = E(X) \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]$$

where $E(X)$ is the mean of the batch size arrivals. (10)

The utilisation of the first service station is

$$U_1(t) = 1 - \exp\{B_1(t)\} \tag{11}$$

The throughput of the service station is

$$Th_1(t) = \mu_1 [1 - \exp\{B_1(t)\}] \tag{12}$$

The average waiting time of a customer in the first queue is

$$W_1(t) = \frac{E(X) \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]}{\mu_1 [1 - \exp\{B_1(t)\}]} \tag{13}$$

The variance of the number of customers in the first queue is

$$V_1(t) = E(X) \left[\frac{\lambda}{\mu_1} \left[(k-1) \left(\frac{1 - e^{-2\mu_1 t}}{2} \right) + (1 - e^{-\mu_1 t}) \right] + \frac{\alpha}{\mu_1^2} \left[(k-1) \left(\frac{2\mu_1 t - 1 + e^{-2\mu_1 t}}{4} \right) + (t\mu_1 - 1 + e^{-\mu_1 t}) \right] \right] \tag{14}$$

The coefficient of variation of the number of customers in the first queue is

$$CV_1(t) = \frac{\sqrt{V_1(t)}}{L_1(t)} \times 100 \tag{15}$$

Taking $Z_1 = 1$ in $P(Z_1, Z_2, t)$, we determine probability generating function for the second queue size to be

$$P(Z_2, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i C_k \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{1 - e^{-[i\mu_2 + (r-i)\mu_1]t}}{i\mu_2 + (r-i)\mu_1} \right] + \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i C_k \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{t[i\mu_2 + (r-i)\mu_1] - 1 + e^{-[i\mu_2 + (r-i)\mu_1]t}}{[i\mu_2 + (r-i)\mu_1]^2} \right] \right\} \tag{16}$$

By expanding $P(Z_2, t)$ and collecting the constant terms, we obtain the probability that the second queue is empty to be

$$P_{0,0}(t) = \exp\{B_2(t)\} \text{ say } B_2(t) = \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} C_k \binom{k}{r} \binom{r}{i} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{1 - e^{-[i\mu_2 + (r-i)\mu_1]t}}{i\mu_2 + (r-i)\mu_1} \right] + \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} C_k \binom{k}{r} \binom{r}{i} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{t[i\mu_2 + (r-i)\mu_1] - 1 + e^{-[i\mu_2 + (r-i)\mu_1]t}}{[i\mu_2 + (r-i)\mu_1]^2} \right] \tag{17}$$

The mean number of customers in the second queue is

$$L_2(t) = E(X) \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right] \tag{18}$$

where $E(X)$ is the mean of the batch size arrivals.

The utilisation of the second service station is

$$U_2(t) = 1 - \exp\{B_2(t)\} \tag{19}$$

The throughput of the service station is

$$Th_2(t) = \mu_2 [1 - \exp\{B_2(t)\}] \tag{20}$$

The average waiting time of a customer in the second queue is

$$W_2(t) = \frac{E(X) \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]}{\mu_2 [1 - \exp\{B_2(t)\}]}$$
(21)

The variance of the number of customers in the second queue is

$$V_2(t) = E(X) \frac{(k-1) \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^2 \left(\lambda \left[\left(\frac{1 - e^{-2\mu_1 t}}{\mu_1} \right) - 4 \left(\frac{1 - e^{-(\mu_2 + \mu_1)t}}{\mu_2 + \mu_1} \right) + \left(\frac{1 - e^{-2\mu_2 t}}{\mu_2} \right) \right] \right.}{+ \alpha \left[\left(\frac{2\mu_1 t - 1 + e^{-2\mu_1 t}}{2\mu_1^2} \right) - 4 \left(\frac{t[\mu_2 + \mu_1] - 1 + e^{-(\mu_2 + \mu_1)t}}{[\mu_2 + \mu_1]^2} \right) \right.}$$

$$+ \left. \left. \left(\frac{2\mu_2 t - 1 + e^{-2\mu_2 t}}{2\mu_2^2} \right) \right] \right) + E(X) \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]$$
(22)

The coefficient of variation of the number of customers in the second queue is

$$CV_2(t) = \frac{\sqrt{V_2(t)}}{L_2(t)} \times 100$$
(23)

The mean number of customers in the queueing system at time t is

$$L(t) = L_1(t) + L_2(t)$$
(24)

where, $L_1(t)$, $L_2(t)$ are as given equations (10) and (18).

4. CHARACTERISTICS OF THE MODEL UNDER UNIFORM BATCH SIZE DISTRIBUTION

The performance measures of the queueing model are highly influenced by the form of the batch size distribution. It is assumed that the number of customers in any arriving module is random and follows a uniform distribution with parameters a and b. Then, the probability mass function of the batch size distributions is $C_k = \frac{1}{b-a+1}$ for $k = a, a + 1, \dots, b$ Q_1 and Q_2 respectively.

The mean number of customers in the batch is

$$E(X) = \frac{a + b}{2}$$

The variance of the batch size is

$$V(X) = \frac{1}{12} [(b - a + 1)^2 - 1]$$

The joint probability generating function for the number of the customers in each queue is

$$\begin{aligned}
 P(Z_1, Z_2, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2-1)\mu_1}{\mu_2-\mu_1} \right]^i \left[(Z_1-1) \right. \right. \\
 \left. \left. + \frac{(Z_2-1)\mu_1}{\mu_2-\mu_1} \right]^{r-i} \left[\frac{1 - e^{-[i\mu_2+(r-i)\mu_1]t}}{i\mu_2+(r-i)\mu_1} \right] \right. \\
 \left. + a \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2-1)\mu_1}{\mu_2-\mu_1} \right]^i \left[(Z_1-1) \right. \right. \\
 \left. \left. + \frac{(Z_2-1)\mu_1}{\mu_2-\mu_1} \right]^{r-i} \left[\frac{t[i\mu_2+(r-i)\mu_1] - 1 + e^{-[i\mu_2+(r-i)\mu_1]t}}{[i\mu_2+(r-i)\mu_1]^2} \right] \right\}, |Z_1| < 1, |Z_2| < 1
 \end{aligned}
 \tag{25}$$

Taking $Z_1 = Z_2 = 0$ in $P(Z_1, Z_2, t)$, we determine the probability generating function for the queueing system is empty to be

$$\begin{aligned}
 P_{00}(t) \\
 = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{\mu_2^{r-i}\mu_1^i}{(\mu_2-\mu_1)^r} \right] \left[\frac{1 - e^{-[i\mu_2+(r-i)\mu_1]t}}{i\mu_2+(r-i)\mu_1} \right] \right. \\
 \left. + a \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{\mu_2^{r-i}\mu_1^i}{(\mu_2-\mu_1)^r} \right] \left[\frac{t[i\mu_2+(r-i)\mu_1] - 1 + e^{-[i\mu_2+(r-i)\mu_1]t}}{[i\mu_2+(r-i)\mu_1]^2} \right] \right\}
 \end{aligned}
 \tag{26}$$

Taking $Z_2 = 1$ in $P(Z_1, Z_2, t)$, we determine probability generating function for the first queue size to be

$$\begin{aligned}
 P(Z_1, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \frac{1}{b-a+1} \binom{k}{r} (Z_1-1)^r \left[\frac{1 - e^{-r\mu_1 t}}{r\mu_1} \right] \right. \\
 \left. + a \sum_{k=1}^{\infty} \sum_{r=1}^k \frac{1}{b-a+1} \binom{k}{r} (Z_1-1)^r \left[\frac{tr\mu_1 - 1 + e^{-r\mu_1 t}}{[r\mu_1]^2} \right] \right\}
 \end{aligned}
 \tag{27}$$

By expanding $P(Z_1, t)$ and collecting the constant terms, we determine the probability that the first queue is empty to be

$$\begin{aligned}
 P_0(t) = \exp\{B_3(t)\} \text{ say} \\
 \text{where } B_3(t) = \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \frac{1}{b-a+1} \binom{k}{r} (-1)^r \left[\frac{1 - e^{-r\mu_1 t}}{r\mu_1} \right] \\
 + a \sum_{k=1}^{\infty} \sum_{r=1}^k \frac{1}{b-a+1} \binom{k}{r} (-1)^r \left[\frac{tr\mu_1 - 1 + e^{-r\mu_1 t}}{[r\mu_1]^2} \right]
 \end{aligned}
 \tag{28}$$

The mean number of customers in the first queue is

$$L_1(t) = \frac{a+b}{2} \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{a}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right] \tag{29}$$

The utilisation of the first service station is

$$U_1(t) = 1 - \exp\{B_3(t)\} \tag{30}$$

The throughput of the service station is

$$Th_1(t) = \mu_1 [1 - \exp\{B_3(t)\}] \tag{31}$$

The average waiting time of a customer in the first queue is

$$W_1(t) = \frac{\frac{a+b}{2} \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{a}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]}{\mu_1 [1 - \exp\{B_3(t)\}]} \tag{32}$$

The variance of the number of customers in the first queue is

$$V_1(t) = \frac{a+b}{2} \left[\frac{\lambda}{\mu_1} \left[(k-1) \left(\frac{1 - e^{-2\mu_1 t}}{2} \right) + (1 - e^{-\mu_1 t}) \right] + \frac{a}{\mu_1^2} \left[(k-1) \left(\frac{2\mu_1 t - 1 + e^{-2\mu_1 t}}{4} \right) + (t\mu_1 - 1 + e^{-\mu_1 t}) \right] \right] \tag{33}$$

The coefficient of variation of the number of customers in the first queue is

$$CV_1(t) = \frac{\sqrt{V_1(t)}}{L_1(t)} \times 100 \tag{34}$$

Taking $Z_1 = 1$ in $P(Z_1, Z_2, t)$, we determine probability generating function for the second queue size to be

$$P(Z_2, t) = \exp \left\{ \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{1 - e^{-[i\mu_2 + (r-i)\mu_1]t}}{i\mu_2 + (r-i)\mu_1} \right] + a \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^i \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{(Z_2 - 1)\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{t[i\mu_2 + (r-i)\mu_1] - 1 + e^{-[i\mu_2 + (r-i)\mu_1]t}}{[i\mu_2 + (r-i)\mu_1]^2} \right] \right\} \tag{35}$$

By expanding $P(Z_2, t)$ and collecting the constant terms, we determine the probability that the second queue is empty to be

$$P_{,0}(t) = \exp\{B_4(t)\} \text{ say}$$

where $B_2(t)$

$$\begin{aligned}
 &= \lambda \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{1 - e^{-[i\mu_2 + (r-i)\mu_1]t}}{i\mu_2 + (r-i)\mu_1} \right] \\
 &+ \alpha \sum_{k=1}^{\infty} \sum_{r=1}^k \sum_{i=0}^r (-1)^{i+r} \frac{1}{b-a+1} \binom{k}{r} \binom{r}{i} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^r \left[\frac{t[i\mu_2 + (r-i)\mu_1] - 1 + e^{-[i\mu_2 + (r-i)\mu_1]t}}{[i\mu_2 + (r-i)\mu_1]^2} \right]
 \end{aligned}
 \tag{36}$$

The mean number of customers in the second queue is

$$L_2(t) = \frac{a+b}{2} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]
 \tag{37}$$

The utilisation of the second service station is

$$U_2(t) = 1 - \exp\{B_2(t)\}
 \tag{38}$$

The throughput of the service station is

$$Th_2(t) = \mu_2 [1 - \exp\{B_2(t)\}]
 \tag{39}$$

The average waiting time of a customer in the second queue is

$$W_2(t) = \frac{\frac{a+b}{2} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]}{\mu_2 [1 - \exp\{B_2(t)\}]}
 \tag{40}$$

The variance of the number of customers in the second queue is

$$\begin{aligned}
 V_2(t) &= \frac{a+b}{2} \frac{(k-1)}{2} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right]^2 \left(\lambda \left[\left(\frac{1 - e^{-2\mu_1 t}}{\mu_1} \right) - 4 \left(\frac{1 - e^{-(\mu_2 + \mu_1)t}}{\mu_2 + \mu_1} \right) + \left(\frac{1 - e^{-2\mu_2 t}}{\mu_2} \right) \right] \right. \\
 &+ \alpha \left[\left(\frac{2\mu_1 t - 1 + e^{-2\mu_1 t}}{2\mu_1^2} \right) - 4 \left(\frac{t[\mu_2 + \mu_1] - 1 + e^{-[\mu_2 + \mu_1]t}}{[\mu_2 + \mu_1]^2} \right) \right. \\
 &+ \left. \left. \left(\frac{2\mu_2 t - 1 + e^{-2\mu_2 t}}{2\mu_2^2} \right) \right] \right) \\
 &+ \frac{a+b}{2} \left[\frac{\mu_1}{\mu_2 - \mu_1} \right] \left[\frac{\lambda}{\mu_1} [1 - e^{-\mu_1 t}] + \frac{\alpha}{\mu_1^2} [t\mu_1 - 1 + e^{-\mu_1 t}] \right]
 \end{aligned}
 \tag{41}$$

The coefficient of variation of the number of customers in the second queue is

$$CV_2(t) = \frac{\sqrt{V_2(t)}}{L_2(t)} \times 100
 \tag{42}$$

The mean number of customers in the queueing system at time t is

$$L(t) = L_1(t) + L_2(t)
 \tag{43}$$

where $L_1(t)$, $L_2(t)$ are as given equations (29),(37).

5. NUMERICAL DEMONSTRATION AND SENSITIVITY ANALYSIS:

In this section, the performance of the proposed queueing model is discussed through a numerical illustration. The customers arrive in batches to the first queue and after getting service through the first server join the second queue, which is serially connected to the first service station. The arrival of the customers follows a compound Poisson process the composite mean arrival rate is $\lambda(t) = \lambda + \alpha t$. Each arriving module represents a batch of customers. The number of customers in each arriving module follows a uniform distribution with parameters (a,b). Because the characteristics of the queueing model are highly sensitive with respect to time, the transient behaviour of the model is studied by computing the performance measures with the following set of values for the model parameters:

$t = 0.05, 0.06, 0.07, 0.08, 0.09$; $a = 2, 3, 4, 5, 6$; $b = 12, 14, 16, 18, 20$; $\lambda = 1, 1.5, 2, 2.5, 3$;
 $\alpha = 0, 1, 1.5, 2, 2.5$; $\mu_1 = 16, 17, 18, 19, 20$; $\mu_2 = 22, 23, 24, 25, 26$;

From the equations of the probability that the system is empty and the emptiness of the marginal queues, the expected number of customers, and the utilisation of servers are computed for different values of the parameters $t, a, b, \lambda, \alpha, \mu_1, \mu_2$, and presented in Table 1. The relationships between the parameters and performance measures are shown in figure 2.

From Table 1, it can be observed that the probability of emptiness of the queueing system and two marginal queues are highly sensitive with respect to time. As time (t) increases from 0.05 to 0.09 the probability that the queue is empty decreases; the probability that the marginal queues is empty decreases; the expected number of customers in each queue and in the system increases; and the utilisation of the service station increases, when all other parameters are fixed.

As the parameter a increases from 2 to 6, the probability that the queue is empty unchanged; the probability that the marginal queues is empty decreases; the expected number of customers in each queue and in the system increases; and the utilisation of the service station increases, when all other parameters are fixed. As the parameter b increases from 12 to 20, the probability that the queue is empty decreases; the probability that the marginal queues is empty decreases; the expected number of customers in each queue and in the system increases; and the utilisation of the service station increases, when all other parameters are fixed.

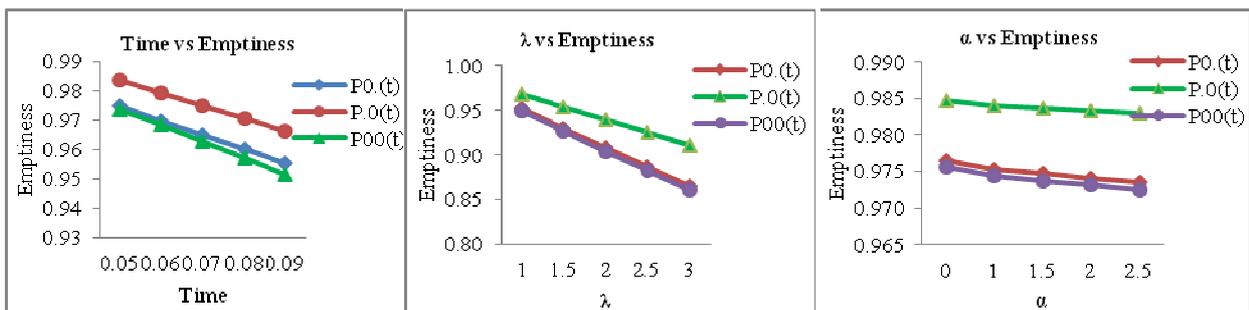
It can be further observed that as the parameter (λ) increases from 1 to 3, the probability that the queue is empty decreases; the probability that the marginal queue is empty decreases; the expected number of customers in each queue and in the system increases; and the utilisation of the service station increases, when all other parameters are fixed. The same patterns hold with respect to the parameter (α). It can also be observed that as the service rate (μ_1) increases from 16 to 20, the probability that the queue is empty remains constant; the probability that the first queue increases, but the second queue decreases; the expected number of customers in the first queue decreases, but the second queue and the system increases; and the utilisation of the service station in the first queue decreases, but the second queue increases, when all other parameters are fixed. It can also be observed that as the service rate (μ_2) increases from 22 to 26, the probability that the queue is empty increases and remains constant; the probability that the second queue increases, but the first queue it remains constant; the expected number of customers in the queue, and the utilisation of the service station in the second queue decreases, but the first queue it remains constant.

Table – 1

Values of $P_0(t)$, $P_{00}(t)$, $L_1(t)$, $L_2(t)$, $L(t)$, $U_1(t)$, $U_2(t)$ for different values of parameters

t	a	b	λ	α	μ_1	μ_2	$P_{00}(t)$	$P_0(t)$	$P_{00}(t)$	$L_1(t)$	$L_2(t)$	$L(t)$	$U_1(t)$	$U_2(t)$
0.05	1	10	0.5	1.5	15	21	0.9747	0.9837	0.9737	0.1049	0.2622	0.3671	0.0253	0.0163
0.06							0.9697	0.9793	0.9683	0.1200	0.3001	0.4201	0.0303	0.0207
0.07							0.9648	0.9749	0.9627	0.1338	0.3346	0.4684	0.0352	0.0251
0.08							0.9600	0.9704	0.9572	0.1465	0.3662	0.5127	0.0400	0.0296
0.09							0.9553	0.9660	0.9516	0.1581	0.3954	0.5535	0.0447	0.0340
	2						0.9740	0.9825	0.9735	0.1144	0.2861	0.4005	0.0260	0.0175
	3						0.9737	0.9815	0.9735	0.1240	0.3099	0.4339	0.0263	0.0185
	4						0.9736	0.9807	0.9735	0.1335	0.3337	0.4672	0.0264	0.0193
	5						0.9736	0.9801	0.9735	0.1430	0.3576	0.5006	0.0264	0.0199
	6						0.9735	0.9795	0.9735	0.1526	0.3814	0.5340	0.0265	0.0205
		12					0.9745	0.9826	0.9737	0.1240	0.3099	0.4339	0.0255	0.0174
		14					0.9744	0.9818	0.9737	0.1430	0.3576	0.5006	0.0256	0.0182
		16					0.9743	0.9811	0.9736	0.1621	0.4052	0.5673	0.0257	0.0189
		18					0.9742	0.9805	0.9736	0.1812	0.4529	0.6341	0.0258	0.0195
		20					0.9741	0.9800	0.9736	0.2002	0.5006	0.7008	0.0259	0.0200
			1				0.9518	0.9685	0.9499	0.2016	0.5040	0.7056	0.0482	0.0315
			1.5				0.9295	0.9537	0.9267	0.2984	0.7459	1.0443	0.0705	0.0463
			2				0.9076	0.9390	0.9041	0.3951	0.9877	1.3828	0.0924	0.0610
			2.5				0.8863	0.9246	0.8820	0.4918	1.2295	1.7213	0.1137	0.0754
			3				0.8655	0.9104	0.8604	0.5886	1.4714	2.0600	0.1345	0.0896
				0			0.9765	0.9846	0.9756	0.0967	0.2418	0.3385	0.0235	0.0154
				1			0.9753	0.9840	0.9743	0.1022	0.2554	0.3576	0.0247	0.0160
				1.5			0.9747	0.9837	0.9737	0.1049	0.2622	0.3671	0.0253	0.0163
				2			0.9741	0.9833	0.9731	0.1076	0.2690	0.3766	0.0259	0.0167
				2.5			0.9735	0.9830	0.9725	0.1103	0.2758	0.3861	0.0265	0.0170
					16		0.9748	0.9833	0.9738	0.1027	0.3286	0.4313	0.0252	0.0167
					17		0.9749	0.9829	0.9738	0.1005	0.4273	0.5278	0.0251	0.0171
					18		0.9751	0.9825	0.9738	0.0985	0.5908	0.6893	0.0249	0.0175
					19		0.9752	0.9822	0.9738	0.0965	0.9163	1.0128	0.0248	0.0178
					20		0.9753	0.9818	0.9738	0.0945	1.8901	1.9846	0.0247	0.0182
					22		0.9747	0.9838	0.9737	0.1049	0.2248	0.3297	0.0253	0.0162
					23		0.9747	0.9839	0.9738	0.1049	0.1967	0.3016	0.0253	0.0161
					24		0.9747	0.9840	0.9738	0.1049	0.1748	0.2797	0.0253	0.0160
					25		0.9747	0.9841	0.9738	0.1049	0.1573	0.2622	0.0253	0.0159
					26		0.9747	0.9841	0.9738	0.1049	0.1430	0.2479	0.0253	0.0159

Figures:



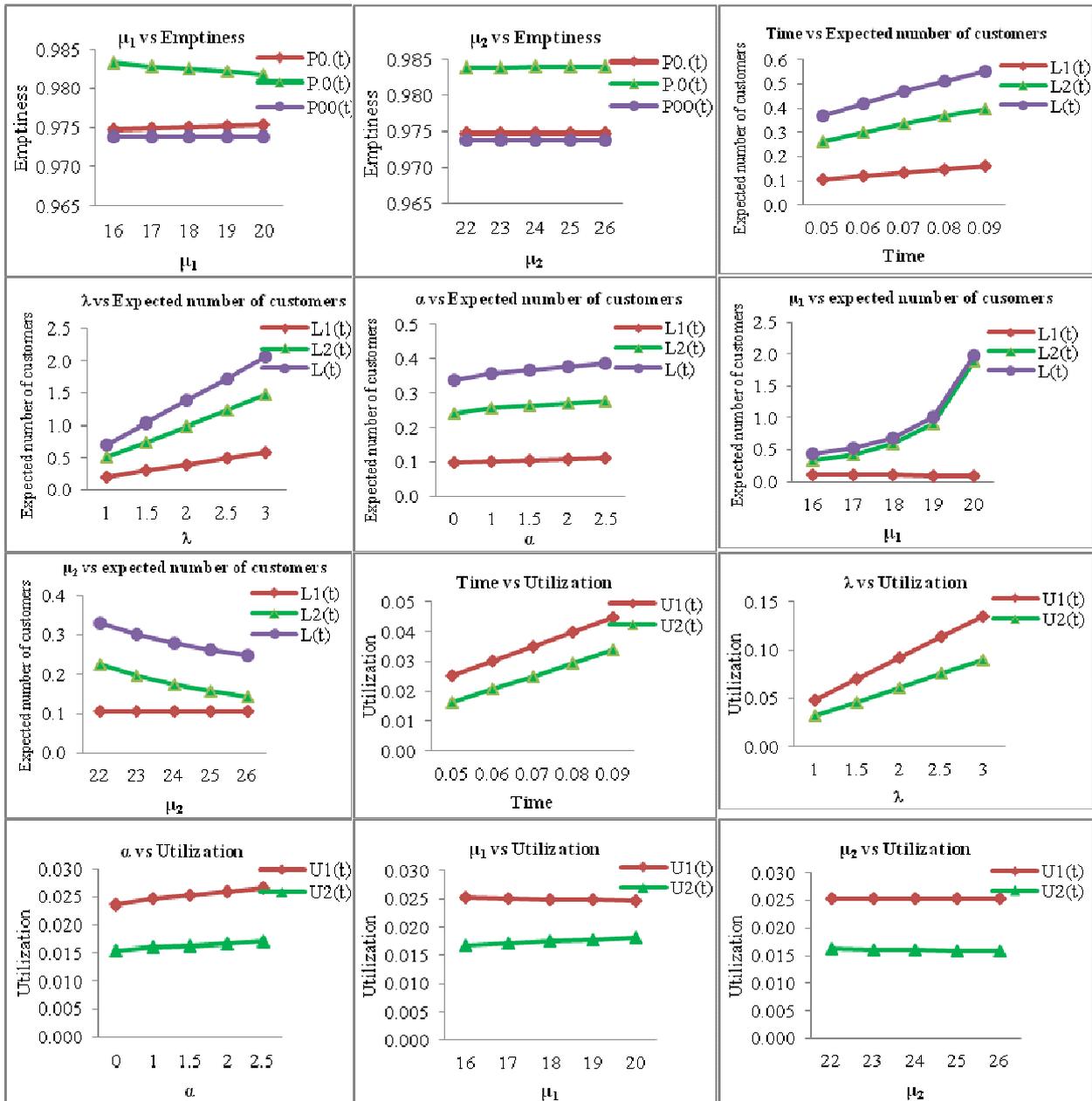


Figure 2: The relationships between the parameters and performance measures.

The throughput of the service stations, the average waiting time, the variance, and the coefficient of variation of number of customers in each queue are computed for different values of t , a , b , λ , α , μ_1 , μ_2 and presented in Table 2. The relationships between parameters and the performance measures are shown in figure 3.

From Table 2 it can be observed that the throughput of the service stations, waiting time of a customer in each queue, variance and coefficient of variation of number of customers in each queue are highly sensitive with respect to time. As time (t) increases from 0.05 to 0.09, the throughput of the service station, and the variance of the number of customers in each queue increases; the average waiting time of a customer, and the coefficient of variation of number of customers in each queue decreases, when all other parameters are fixed. As the parameter a increases from 2 to 6, the throughput of the service station, the waiting time of a customer in each queue, and the variance of the number of customers in each queue increases, but the coefficient of variation of the number of customer in each queue decreases. The same patterns hold with respect to the parameter b .

Table – 2: Values of $Th_1(t)$, $Th_2(t)$, $W_1(t)$, $W_2(t)$, $V_1(t)$, $V_2(t)$, $CV_1(t)$, $CV_2(t)$ for different values of parameter

t	a	b	λ	α	μ_1	μ_2	$Th_1(t)$	$Th_2(t)$	$W_1(t)$	$W_2(t)$	$V_1(t)$	$V_2(t)$	$CV_1(t)$	$CV_2(t)$
0.05	1	10	0.5	1.5	15	21	0.3791	0.3430	0.2767	0.7644	0.5719	0.3074	720.9158	211.4556
0.06							0.4540	0.4343	0.2644	0.6910	0.6322	0.3624	662.5917	200.5987
0.07							0.5278	0.5272	0.2536	0.6347	0.6837	0.4141	617.9832	192.3210
0.08							0.6001	0.6209	0.2441	0.5898	0.7286	0.4625	582.6489	185.7110
0.09							0.6706	0.7150	0.2358	0.5529	0.7684	0.5073	554.4493	180.1340
	2						0.3896	0.3680	0.2937	0.7773	0.6334	0.3362	695.6856	202.6660
	3						0.3939	0.3883	0.3147	0.7980	0.7042	0.3660	676.7470	195.2176
	4						0.3958	0.4050	0.3373	0.8239	0.7845	0.3967	663.4607	188.7448
	5						0.3967	0.4189	0.3605	0.8537	0.8743	0.4282	653.8742	182.9894
	6						0.3972	0.4304	0.3841	0.8861	0.9734	0.4607	646.5337	177.9625
		12					0.3822	0.3652	0.3244	0.8485	0.7986	0.3751	720.9922	197.6296
		14					0.3844	0.3831	0.3721	0.9334	1.0630	0.4465	720.8400	186.8587
		16					0.3861	0.3977	0.4199	1.0190	1.3651	0.5215	720.7742	178.2205
		18					0.3874	0.4099	0.4677	1.1050	1.7050	0.6002	720.6808	171.0589
		20					0.3884	0.4201	0.5155	1.1915	2.0826	0.6825	720.6162	165.0291
			1				0.7226	0.6605	0.2790	0.7631	1.0959	0.5922	519.2720	152.6876
			1.5				1.0581	0.9731	0.2820	0.7665	1.6200	0.8771	426.5389	125.5579
			2				1.3857	1.2809	0.2851	0.7711	2.1440	1.1619	370.6000	109.1338
			2.5				1.7056	1.5840	0.2884	0.7762	2.6680	1.4468	332.1272	97.8308
			3				2.0179	1.8824	0.2917	0.7816	3.1920	1.7317	303.5367	89.4346
				0			0.3525	0.3228	0.2745	0.7493	0.5240	0.2849	748.5816	220.7445
				1			0.3702	0.3363	0.2760	0.7595	0.5560	0.2999	729.6028	214.4210
				1.5			0.3791	0.3430	0.2767	0.7644	0.5719	0.3074	720.9158	211.4556
				2			0.3879	0.3498	0.2774	0.7691	0.5879	0.3149	712.5896	208.6095
				2.5			0.3968	0.3565	0.2781	0.7735	0.6039	0.3224	704.5422	205.8748
					16		0.4026	0.3515	0.2550	0.9348	0.5530	0.3782	724.0893	187.1515
					17		0.4259	0.3593	0.2361	1.1893	0.5350	0.4815	727.7980	162.3922
					18		0.4488	0.3665	0.2194	1.6122	0.5179	0.6495	730.6119	136.4109
					19		0.4715	0.3731	0.2046	2.4560	0.5017	0.9796	733.9978	108.0157
					20		0.4939	0.3816	0.1914	4.9535	0.4863	1.9579	737.9388	74.0305
						22	0.3791	0.3572	0.2767	0.6291	0.5719	0.2685	720.9158	230.5026
						23	0.3791	0.3713	0.2767	0.5297	0.5719	0.2390	720.9158	248.5390
						24	0.3791	0.3851	0.2767	0.4540	0.5719	0.2159	720.9158	265.8183
						25	0.3791	0.3987	0.2767	0.3946	0.5719	0.1971	720.9158	282.2374
						26	0.3791	0.4121	0.2767	0.3470	0.5719	0.1816	720.9158	298.0039

It can be further observed that as the parameter (λ) increases from 1 to 3, the throughput of the service station, the waiting time of the customer in each queue, and the variance of the number of customers in each queue increases, whereas the coefficient of variation of the number of customers in each queue decreases. The same patterns hold with respect to the parameter (α).

It can also be observed that as the service rate (μ_1) increases from 16 to 20, the throughput of the service stations increases; the average waiting time, and the variance of the number of customers in the first queue decreases, but the second queues increases; and the coefficient of variation of the number of customers in the first queue increases, but the second queue decreases. It can also be observed that as the service rate (μ_2) increases from 22 to 26, the throughput of the second service station increases, but the first service station remains constant; the average waiting time, and the variance of the number of customers in the second queue

decreases, but the first queue remains constant; and the coefficient of variation of the number of customers in the second queue increases, but the first queue remains constant.

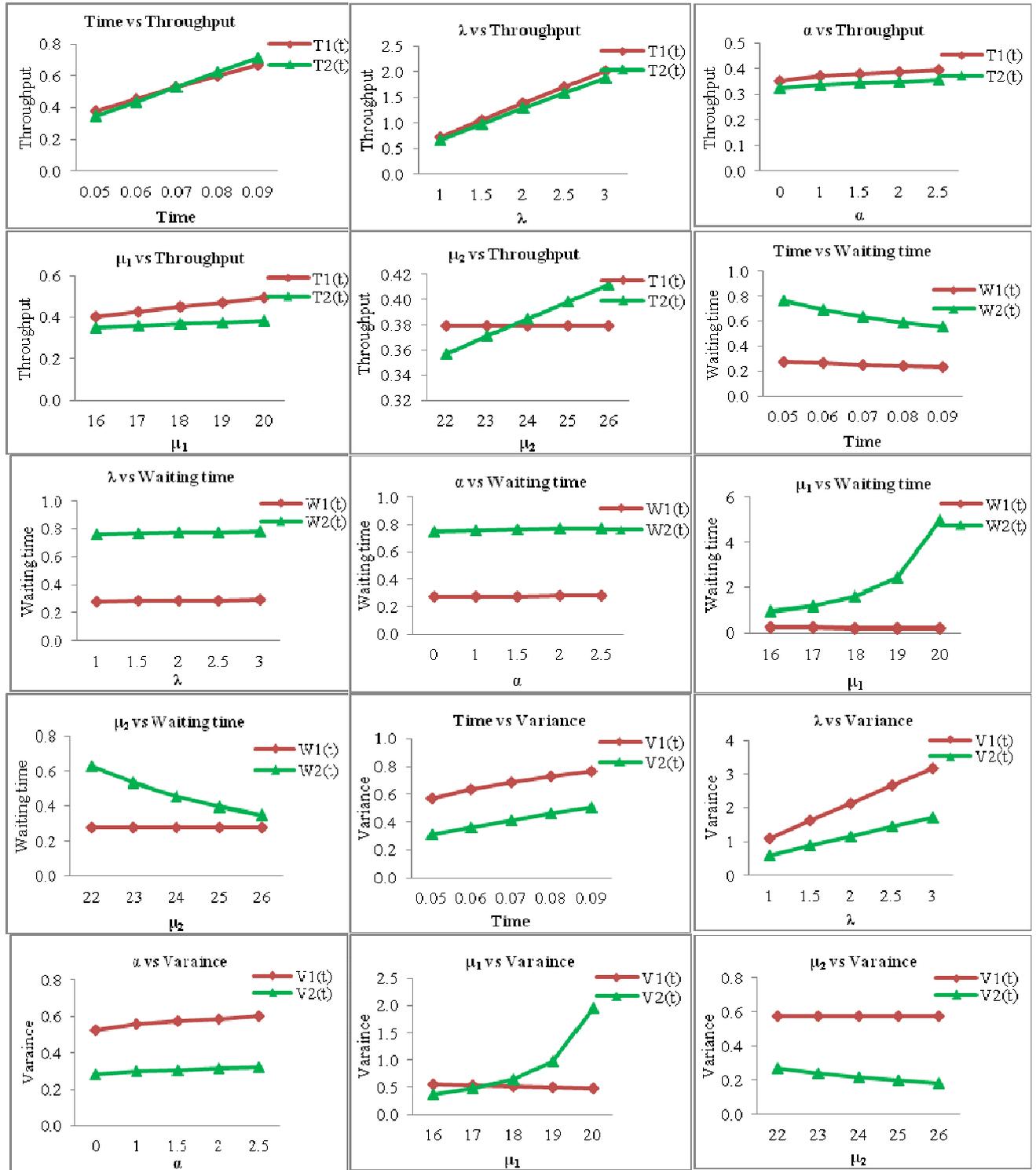


Figure 3: The relationships between the parameters and performance measures.

From this analysis it can also be observed that the bulk size distribution parameters and composite mean arrival significantly influences the system performance measures.

6. COMPARATIVE STUDY

A comparative study of the developed model with that of homogeneous compound Poisson arrivals is carried by taking $\alpha = 0$ in the model and different values of t . Table 3 shows the points study of having models with homogeneous and non-homogeneous compound Poisson arrivals.

Table 3: Comparative study of models with non- homogeneous and homogeneous Poisson arrivals

t	Parameters Measured	$\alpha = 1$	$\alpha = 0$	Difference	Percentage of Variation
0.05	$L_1(t)$	0.1022	0.0967	0.0055	5.6877
	$L_2(t)$	0.2554	0.2418	0.0136	5.6245
	$U_1(t)$	0.0247	0.0235	0.0012	5.1064
	$U_2(t)$	0.0160	0.0154	0.0006	3.8961
	$T_1(t)$	0.3702	0.3525	0.0177	5.0213
	$T_2(t)$	0.3363	0.3228	0.0135	4.1822
	$W_1(t)$	0.2760	0.2745	0.0015	0.5464
	$W_2(t)$	0.7595	0.7493	0.0102	1.3613
0.06	$L_1(t)$	0.1163	0.1088	0.0075	6.8934
	$L_2(t)$	0.2907	0.2720	0.0187	6.8750
	$U_1(t)$	0.0294	0.0277	0.0017	6.1372
	$U_2(t)$	0.0202	0.0192	0.0010	5.2083
	$T_1(t)$	0.4414	0.4162	0.0252	6.0548
	$T_2(t)$	0.4239	0.4033	0.0206	5.1079
	$W_1(t)$	0.2635	0.2614	0.0021	0.8034
	$W_2(t)$	0.6858	0.6744	0.0114	1.6904
0.07	$L_1(t)$	0.1290	0.1192	0.0098	8.2215
	$L_2(t)$	0.3224	0.2979	0.0245	8.2242
	$U_1(t)$	0.0341	0.0318	0.0023	7.2327
	$U_2(t)$	0.0244	0.0230	0.0014	6.0870
	$Th_1(t)$	0.5109	0.4770	0.0339	7.1069
	$Th_2(t)$	0.5125	0.4832	0.0293	6.0637
	$W_1(t)$	0.2524	0.2498	0.0026	1.0408
	$W_2(t)$	0.6290	0.6166	0.0124	2.0110
0.08	$L_1(t)$	0.1404	0.1281	0.0123	9.6019
	$L_2(t)$	0.3509	0.3203	0.0306	9.5535
	$U_1(t)$	0.0386	0.0356	0.0030	8.4270
	$U_2(t)$	0.0286	0.0267	0.0019	7.1161
	$T_1(t)$	0.5784	0.5347	0.0437	8.1728
	$T_2(t)$	0.6012	0.5617	0.0395	7.0322
	$W_1(t)$	0.2427	0.2396	0.0031	1.2938
	$W_2(t)$	0.5837	0.5702	0.0135	2.3676
0.09	$L_1(t)$	0.1507	0.1358	0.0149	10.9720
	$L_2(t)$	0.3767	0.3395	0.0372	10.9573
	$U_1(t)$	0.0429	0.0393	0.0036	9.1603
	$U_2(t)$	0.0328	0.0304	0.0024	7.8947
	$T_1(t)$	0.6434	0.5890	0.0544	9.2360
	$T_2(t)$	0.6895	0.6384	0.0511	8.0044
	$W_1(t)$	0.2342	0.2306	0.0036	1.5611
	$W_2(t)$	0.5464	0.5318	0.0146	2.7454

From the table 3 it can also be observed that as time increases, the percentage variation of the performance measures between the models also increases. The model with non-homogeneous compound Poisson arrivals has higher utilisation than the model with homogeneous compound Poisson arrivals. It can also be observed that the assumption of non-homogenous compound Poisson arrivals has a significant influence on all the performance measures of the queueing model. Time also has a significant effect on the system performance measures, and this model can predict the performance measures more accurately. This model also includes some of the earlier models as particular cases.

7. CONCLUSION:

This paper addresses the development and analysis of a two-node tandem queueing model with non-homogeneous bulk arrivals and a state-dependent service rate. It is also assumed that the two queues are connected in tandem and after getting service from the first queue the customer join the second queue. It is also assumed that the customers arrive in bulk groups of random size depending on time. The arrival process of the queue is characterised by a non-homogeneous compound Poisson process. The explicit expressions for the system characteristics, such as the average number of customers in each queue, the probability that the queue is empty, the average waiting time of a customer in the queue, and the throughput of the service station are derived explicitly. The sensitivity of the model revealed that the bulk size distribution parameters have a significant influence on the system performance measures. By regulating the bulk size distribution parameters the congestion in queues and the mean delay can be reduced. A comparative study of the developed model with a model using homogeneous compound Poisson arrivals revealed that time has a significant effect on system performance measures and the performance measures can be predicted more accurately and realistically using the developed model. This model also includes several of the earlier existing models as particular cases for specific values of the parameters. This model can also be extended by obtaining the optimal values of the model parameters under cost considerations, which will be pursued elsewhere.

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